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18.175 Theory of Probability Fall 2008

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## Section 2

# Random variables and their properties. Expectation.

Let  $(\Omega, \mathcal{A}, \mathbb{P})$  be a probability space and  $(\mathcal{S}, \mathcal{B})$  be a measurable space where  $\mathcal{B}$  is a  $\sigma$ -algebra of subsets of  $\mathcal{S}$ . A random variable  $X : \Omega \to \mathcal{S}$  is a measurable function, i.e.

$$B \in \mathcal{B} \Longrightarrow X^{-1}(B) \in \mathcal{A}.$$

When  $S = \mathbb{R}$  we will usually consider a  $\sigma$ -algebra  $\mathcal{B}$  of Borel measurable sets generated by sets  $\bigcup_{i \leq n} (a_i, b_i]$  (or, equivalently, generated by sets  $(a_i, b_i)$  or by open sets).

**Lemma 3**  $X: \Omega \to \mathbb{R}$  is a random variable iff for all  $t \in \mathbb{R}$ 

$${X \le t} := {\omega \in \Omega : X(\omega) \in (-\infty, t]} \in \mathcal{A}.$$

**Proof.** Only  $\Leftarrow$  direction requires proof. We will prove that

$$\mathcal{D} = \{ D \subseteq \mathbb{R} : X^{-1}(D) \in \mathcal{A} \}$$

is a  $\sigma$ -algebra. Since sets  $(-\infty, t] \in \mathcal{D}$  this will imply that  $\mathcal{B} \subseteq \mathcal{D}$ . The result follows simply because taking pre-image preserves set operations. For example, if we consider a sequence  $D_i \in \mathcal{D}$  for  $i \geq 1$  then

$$X^{-1}\left(\bigcup_{i>1}D_i\right) = \bigcup_{i>1}X^{-1}(D_i) \in \mathcal{A}$$

because  $X^{-1}(D_i) \in \mathcal{A}$  and  $\mathcal{A}$  is a  $\sigma$ -algebra. Therefore,  $\bigcup_{i \geq 1} D_i \in \mathcal{D}$ . Other properties can be checked similarly, so  $\mathcal{D}$  is a  $\sigma$ -algebra.

Let us define a measure  $\mathbb{P}_X$  on  $\mathcal{B}$  by  $\mathbb{P}_X = \mathbb{P} \circ X^{-1}$ , i.e. for  $B \in \mathcal{B}$ ,

$$\mathbb{P}_X(B) = \mathbb{P}(X \in B) = \mathbb{P}(X^{-1}(B)) = \mathbb{P} \circ X^{-1}(B).$$

 $(S, \mathcal{B}, \mathbb{P}_X)$  is called the *sample space* of a random variable X and  $\mathbb{P}_X$  is called *the law* of X. Clearly, on this space a random variable  $\xi : S \to S$  defined by the identity  $\xi(s) = s$  has the same law as X.

When  $S = \mathbb{R}$ , a function  $F(t) = \mathbb{P}(X \le t)$  is called the cumulative distribution function (c.d.f.) of X.

Lemma 4 F is a c.d.f. of some r.v. X iff

- 1.  $0 \le F(t) \le 1$ ,
- 2. F is non-decreasing, right-continuous,

3. 
$$\lim_{t \to -\infty} F(t) = 0$$
,  $\lim_{t \to +\infty} F(t) = 1$ .

**Proof.** The fact that any c.d.f. satisfies properties 1 - 3 is obvious. Let us show that F which satisfies properties 1 - 3 is a c.d.f. of some r.v. X. Consider algebra A consisting of sets  $\bigcup_{i < n} (a_i, b_i]$  for disjoint intervals and for all  $n \geq 1$ . Let us define a function  $\mathbb{P}$  on A by

$$\mathbb{P}\Big(\bigcup_{i < n} (a_i, b_i]\Big) = \sum_{i < n} \big(F(a_i) - F(b_i)\big).$$

One can show that  $\mathbb{P}$  is countably additive on A. Then, by Caratheodory extension Theorem 1,  $\mathbb{P}$  extends uniquely to a measure  $\mathbb{P}$  on  $\sigma(A) = \mathcal{B}$  - Borel measurable sets. This means that  $(\mathbb{R}, \mathcal{B}, \mathbb{P})$  is a probability space and, clearly, random variable  $X: \mathbb{R} \to \mathbb{R}$  defined by X(x) = x has c.d.f.  $\mathbb{P}(X \le t) = F(t)$ . Below we will sometimes abuse the notations and let F denote both c.d.f. and probability measure  $\mathbb{P}$ .

**Alternative proof.** Consider a probability space ([0,1],  $\mathcal{B}$ ,  $\lambda$ ), where  $\lambda$  is the Lebesgue measure. Define r.v.  $X:[0,1]\to\mathbb{R}$  by the quantile transformation

$$X(t) = \inf\{x \in \mathbb{R}, F(x) \ge t\}.$$

The c.d.f. of X is  $\lambda(t:X(t) \leq a) = F(a)$  since

$$X(t) \le a \iff \inf\{x : F(x) \ge t\} \le a \iff \exists a_n \to a, F(a_n) \ge t \iff F(a) \ge t.$$

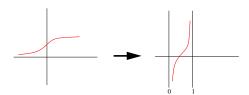


Figure 2.1: A random variable defined by quantile transformation.

**Definition.** Given a probability space  $(\Omega, \mathcal{A}, \mathbb{P})$  and a r.v.  $X : \Omega \to \mathcal{S}$  let  $\sigma(X)$  be a  $\sigma$ -algebra generated by a collection of sets  $\{X^{-1}(B): B \in \mathcal{B}\}$ . Clearly,  $\sigma(X) \subseteq \mathcal{A}$ . Moreover, the above collection of sets is itself a  $\sigma$ -algebra. Indeed, consider a sequence  $A_i = X^{-1}(B_i)$  for some  $B_i \in \mathcal{B}$ . Then

$$\bigcup_{i \ge 1} A_i = \bigcup_{i \ge 1} X^{-1}(B_i) = X^{-1} \left( \bigcup_{i \ge 1} B_i \right) = X^{-1}(B)$$

where  $B \in \bigcup_{i \geq 1} B_i \in \mathcal{B}$ .  $\sigma(X)$  is called the  $\sigma$ -algebra generated by a r.v. X.

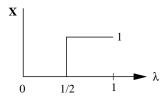


Figure 2.2:  $\sigma(X)$  generated by X.

**Example.** Consider a r.v. defined in figure 2.2. We have  $\mathbb{P}(X=0)=\frac{1}{2}, \mathbb{P}(X=1)=\frac{1}{2}$  and

$$\sigma(X) = \left\{\emptyset, \left[0, \frac{1}{2}\right], \left(\frac{1}{2}, 1\right], [0, 1]\right\}.$$

**Lemma 5** Consider a probability space  $(\Omega, \mathcal{A}, \mathbb{P})$ , a measurable space  $(\mathcal{S}, \mathcal{B})$  and random variables  $X : \Omega \to \mathcal{S}$  and  $Y : \Omega \to \mathbb{R}$ . Then the following are equivalent:

- 1. Y = g(X) for some (Borel) measurable function  $g: S \to \mathbb{R}$ .
- 2.  $Y: \Omega \to \mathbb{R}$  is measurable on  $(\Omega, \sigma(X))$ , i.e. with respect to the  $\sigma$ -algebra generated by X.

**Remark.** It should be obvious from the proof that  $\mathbb{R}$  can be replaced by any separable metric space.

**Proof.** The fact that 1 implies 2 is obvious since for any Borel set  $B \subseteq \mathbb{R}$  the set  $B' := g^{-1}(B) \in \mathcal{B}$  and, therefore,

$${Y = g(X) \in B} = {X \in g^{-1}(B) = B'} = X^{-1}(B') \in \sigma(X).$$

Let us show that 2 implies 1. For all integer n and k consider sets

$$A_{n,k} = \left\{ \omega : Y(\omega) \in \left[ \frac{k}{2^n}, \frac{k+1}{2^n} \right] \right\} = Y^{-1} \left( \left[ \frac{k}{2^n}, \frac{k+1}{2^n} \right] \right).$$

By 2,  $A_{n,k} \in \sigma(X) = \{X^{-1}(B) : B \in \mathcal{B}\}$  and, therefore,  $A_{n,k} = X^{-1}(B_{n,k})$  for some  $B_{n,k} \in \mathcal{B}$ . Let us consider a function

$$g_n(X) = \sum_{k \in \mathbb{Z}} \frac{k}{2^n} I(X \in B_{n,k}).$$

By construction,  $|Y - g_n(X)| \leq \frac{1}{2^n}$  since

$$Y(\omega) \in \left[\frac{k}{2^n}, \frac{k+1}{2^n}\right) \iff X(\omega) \in B_{n,k} \iff g_n(X(\omega)) = \frac{k}{2^n}$$

It is easy to see that  $g_n(x) \leq g_{n+1}(x)$  and, therefore,  $g(x) = \lim_{n \to \infty} g_n(x)$  is a measurable function on  $(\mathcal{S}, \mathcal{B})$  and, clearly, Y = g(X).

#### Discrete random variables.

A r.v.  $X : \Omega \to \mathcal{S}$  is called discrete if  $\mathbb{P}_X(\{S_i\}_{i\geq 1}) = 1$  for some sequence  $S_i \in \mathcal{S}$ .

#### Absolutely continuous random variables.

On a measure space (S, B), a measure  $\mathbb{P}$  is called absolutely continuous w.r.t. a measure  $\lambda$  if

$$\forall B \in \mathcal{B}, \lambda(B) = 0 \Longrightarrow \mathbb{P}(B) = 0.$$

The following is a well known result from measure theory.

**Theorem 2** (Radon-Nikodym) If  $\mathbb{P}$  and  $\lambda$  are sigma-finite and  $\mathbb{P}$  is absolutely continuous w.r.t.  $\lambda$  then there exists a Radon-Nikodym derivative  $f \geq 0$  such that for all  $B \in \mathcal{B}$ 

$$\mathbb{P}(B) = \int_{B} f(s) d\lambda(s).$$

f is uniquely defined up to a  $\lambda$ -null sets.

In a typical setting of  $S = \mathbb{R}^k$ , a probability measure  $\mathbb{P}$  and Lebesgue's measure  $\lambda$ , f is called the *density* of the distribution  $\mathbb{P}$ .

### Independence.

Consider a probability space  $(\Omega, \mathcal{C}, \mathbb{P})$  and two  $\sigma$ -algebras  $\mathcal{A}, \mathcal{B} \subseteq \mathcal{C}$ .  $\mathcal{A}$  and  $\mathcal{B}$  are called *independent* if

$$\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B)$$
 for all  $A \in \mathcal{A}, B \in \mathcal{B}$ .

 $\sigma$ -algebras  $\mathcal{A}_i \subseteq \mathcal{C}$  for  $i \leq n$  are independent if

$$\mathbb{P}(A_1 \cap \dots \cap A_n) = \prod_{i \le n} \mathbb{P}(A_i) \text{ for all } A_i \in \mathcal{A}_i.$$

 $\sigma$ -algebras  $A_i \subseteq \mathcal{C}$  for  $i \leq n$  are pairwise independent if

$$\mathbb{P}(A_i \cap A_j) = \mathbb{P}(A_i)\mathbb{P}(A_j)$$
 for all  $A_i \in \mathcal{A}_i, A_j \in \mathcal{A}_j, i \neq j$ .

Random variables  $X_i: \Omega \to \mathcal{S}$  for  $i \leq n$  are (pairwise) independent if  $\sigma$ -algebras  $\sigma(X_i), i \leq n$  are (pairwise) independent which is just another convenient way to state the familiar

$$\mathbb{P}(X_1 \in B_1, \dots, X_n \in B_n) = \mathbb{P}(X_1 \in B_1) \times \dots \times \mathbb{P}(X_n \in B_n)$$

for any events  $B_1, \ldots, B_n \in \mathcal{B}$ .

**Example.** Consider a regular tetrahedron die, Figure 2.3, with red, green and blue sides and a red-greenblue base. If we roll this die then indicators of different colors provide an example of pairwise independent r.v.s that are not independent since

$$\mathbb{P}(r) = \mathbb{P}(b) = \mathbb{P}(g) = \frac{1}{2} \text{ and } \mathbb{P}(rb) = \mathbb{P}(rg) = \mathbb{P}(bg) = \frac{1}{4}$$

but

$$\mathbb{P}(rbg) = \frac{1}{4} \neq \mathbb{P}(r)\mathbb{P}(b)\mathbb{P}(g) = \left(\frac{1}{2}\right)^3.$$

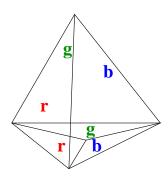


Figure 2.3: Pairwise independent but not independent r.v.s.

Independence of  $\sigma$ -algebras can be checked on generating algebras:

**Lemma 6** If algebras  $A_i$ ,  $i \leq n$  are independent then  $\sigma$ -algebras  $\sigma(A_i)$  are independent.

**Proof.** Obvious by Approximation Lemma 2.

**Lemma 7** Consider  $r.v.s X_i : \Omega \to \mathbb{R}$  on a probability space  $(\Omega, \mathcal{A}, \mathbb{P})$ .

1.  $X_i$ 's are independent iff

$$\mathbb{P}(X_1 \le t_1, \dots, X_n \le t_n) = \mathbb{P}(X_1 \le t_1) \times \dots \times \mathbb{P}(X_n \le t_n). \tag{2.0.1}$$

2. If the laws of  $X_i$ 's have densities  $f_i(x)$  then  $X_i$ 's are independent iff a joint density exists and

$$f(x_1, ..., x_n) = \prod f_i(x_i).$$

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**Proof.** 1 is obvious by Lemma 6 because (2.0.1) implies the same equality for intervals

$$\mathbb{P}(X_1 \in (a_1, b_1], \dots, X_n \in (a_n, b_n]) = \mathbb{P}(X_1 \in (a_1, b_1]) \times \dots \times \mathbb{P}(X_n \in (a_n, b_n])$$

and, therefore, for finite union of disjoint such intervals. To check this for intervals (for example, for n = 2) we can write  $\mathbb{P}(a_1 < X_1 \le b_1, a_2 < X_n \le b_2)$  as

$$\mathbb{P}(X_1 \leq b_1, X_2 \leq b_2) - \mathbb{P}(X_1 \leq a_1, X_2 \leq b_2) - \mathbb{P}(X_1 \leq b_1, X_2 \leq a_2) + \mathbb{P}(X_1 \leq a_1, X_2 \leq a_2)$$

$$= \mathbb{P}(X_1 \leq b_1) \mathbb{P}(X_2 \leq b_2) - \mathbb{P}(X_1 \leq a_1) \mathbb{P}(X_2 \leq b_2) - \mathbb{P}(X_1 \leq b_1) \mathbb{P}(X_2 \leq a_2) + \mathbb{P}(X_1 \leq a_1) \mathbb{P}(X_2 \leq a_2)$$

$$= (\mathbb{P}(X_1 \leq b_1) - \mathbb{P}(X_1 \leq a_1)) (\mathbb{P}(X_2 \leq b_2) - \mathbb{P}(X_2 \leq a_2)) = \mathbb{P}(a_1 < X_1 \leq b_1) \mathbb{P}(a_2 < X_2 \leq b_2).$$

To prove 2 we start with " $\Leftarrow$ ".

$$\begin{split} \mathbb{P}(\cap\{X_i\in A_i\}) &= \mathbb{P}(\boldsymbol{X}\in A_1\times\cdots\times A_n) = \int_{A_1\times\cdots\times A_n} \prod f_i(x_i)d\boldsymbol{x} \\ &= \prod \int_{A_i} f_i(x_i)dx_i \text{ {by Fubini's Theorem}}\} = \prod_{i\leq n} \mathbb{P}(X_i\in A_i). \end{split}$$

Next, we prove "\improx". First of all, by independence,

$$\mathbb{P}(\boldsymbol{X} \in A_1 \times \cdots \times A_n) = \prod \mathbb{P}(X_i \in A_i) \stackrel{\text{Fubini}}{=} \int_{A_1 \times \cdots \times A_n} \prod f_i(x_i) d\boldsymbol{x}.$$

Therefore, the same equality holds for sets in algebra A that consists of finite unions of disjoint sets  $A_1 \times \cdots \times A_n$ , i.e.

$$\mathbb{P}(\boldsymbol{X} \in B) = \int_{B} \prod f_{i}(x_{i}) d\boldsymbol{x} \text{ for } B \in A.$$

Both  $\mathbb{P}(X \in B)$ ,  $\int_{B} \prod f_{i}(x_{i}) dx$  are countably additive on A and finite,

$$\mathbb{P}(\mathbb{R}^n) = \int_{\mathbb{R}^n} \prod f_i(x_i) d\mathbf{x} = 1.$$

By the Caratheodory extension Theorem 1, they extend uniquely to all Borel sets  $\mathcal{B} = \sigma(A)$ , so

$$\mathbb{P}(B) = \int_{B} \prod f_{i}(x_{i}) d\boldsymbol{x} \text{ for } B \in \mathcal{B}.$$

**Expectation.** If  $X:\Omega\to\mathbb{R}$  is a random variable on  $(\Omega,\mathcal{A},\mathbb{P})$  then expectation of X is defined as

$$\mathbb{E}X = \int_{\Omega} X(\omega) d\mathbb{P}(\omega).$$

In other words, expectation is just another term for the integral with respect to a probability measure and, as a result, expectation has all the usual properties of the integrals. Let us emphasize some of them.

**Lemma 8** 1. If F is the c.d.f. of X then for any measurable function  $g: \mathbb{R} \to \mathbb{R}$ ,

$$\mathbb{E}g(x) = \int_{\mathbb{R}} g(x)dF(x).$$

2. If X is discrete, i.e.  $\mathbb{P}(X \in \{x_i\}_{i\geq 1}) = 1$ , then

$$\mathbb{E}X = \sum_{i \ge 1} x_i \mathbb{P}(X = x_i).$$

3. If  $X: \Omega \to \mathbb{R}^k$  has a density f(x) on  $\mathbb{R}^k$  and  $g: \mathbb{R}^k \to \mathbb{R}$  then

$$\mathbb{E}g(X) = \int g(x)f(x)dx.$$

**Proof.** All these properties follow by making a change of variables  $x = X(\omega)$  or  $\omega = X^{-1}(x)$ , i.e.

$$\mathbb{E} g(X) = \int_{\Omega} g(X(\omega)) d\mathbb{P}(\omega) = \int_{\mathbb{R}} g(x) d\mathbb{P} \circ X^{-1}(x) = \int_{\mathbb{R}} g(x) d\mathbb{P}_X(x),$$

where  $\mathbb{P}_X = \mathbb{P} \circ X^{-1}$  is the law of X. Another way to see this would be to start with indicator functions of sets  $g(x) = I(x \in B)$  for which

$$\mathbb{E}g(X) = \mathbb{P}(X \in B) = \mathbb{P}_X(B) = \int_{\mathbb{R}} I(x \in B) d\mathbb{P}_X(x)$$

and, therefore, the same is true for simple step functions

$$g(x) = \sum_{i > n} w_i \mathbf{I}(x \in B_i)$$

for disjoint  $B_i$ . By approximation, this is true for any measurable functions.